

$$(a) L(\lambda) = \prod_{i=1}^n \{2\lambda X_i e^{-\lambda X_i^2}\} = (2\lambda)^n \left(\prod_{i=1}^n X_i\right) e^{-\lambda \sum_{i=1}^n X_i^2}$$

$$\Rightarrow \ell(\lambda) = n \log \lambda + n \log 2 + \sum_{i=1}^n \log X_i - \lambda \sum_{i=1}^n X_i^2$$

$$\Rightarrow \ell'(\lambda) = \frac{n}{\lambda} - \sum_{i=1}^n X_i^2 = 0 \Leftrightarrow \hat{\lambda} = \frac{n}{T}, \text{ όπου } T = \sum_{i=1}^n X_i^2$$

$$\text{και } \ell''(\lambda) = -\frac{n}{\lambda^2} < 0 \quad \forall \lambda \Rightarrow \hat{\lambda} = \frac{n}{T}$$

Τώρα, $Y_i = X_i^2 \sim \text{Exp}(\lambda)$ (επειδή $\lambda > 0$) και X_1, \dots, X_n ανεξ.

$$\Rightarrow T \sim \chi^2(n, \lambda)$$

$$\text{οπότε: } E_{\lambda} \hat{\lambda} = n E_{\lambda} T^{-1} = n \frac{\Gamma(n-1)}{\Gamma(n) \lambda^{-1}} = \frac{n}{n-1} \lambda$$

$$\Rightarrow b(\lambda | \hat{\lambda}_n) = E_{\lambda} \hat{\lambda} - \lambda = \frac{\lambda}{n-1} \neq 0, \text{ δηλ. } \hat{\lambda} \text{ όχι απρόσβ.}$$

(β) Για $c = \frac{n-1}{n}$, $\hat{\theta}_n = c \hat{\lambda}_n = \frac{n-1}{n} \hat{\lambda}_n$ είναι απρόσβ. στην

$$(γ) E(\hat{\lambda})^2 = n^2 E T^{-2} = n^2 \frac{\Gamma(n-2)}{\Gamma(n) \lambda^{-2}} = \frac{n^2}{(n-1)(n-2)} \lambda^2$$

$$\text{Var } \hat{\lambda} = \frac{n^2}{(n-1)(n-2)} \lambda^2 - \frac{n^2}{(n-1)^2} \lambda^2 = \frac{n^2}{(n-1)^2 (n-2)} \lambda^2$$

$$\Rightarrow \bar{b}(\lambda | \hat{\lambda}) = \frac{b(\lambda | \hat{\lambda})}{\sqrt{\text{Var}_{\lambda} \hat{\lambda}}} = \frac{\lambda/(n-1)}{\lambda/(n-1)} = \frac{\sqrt{n-2}}{n} \xrightarrow{n \rightarrow \infty} 0$$

$$(δ) P_{\lambda}(|\hat{\lambda} - \lambda| > \varepsilon) \leq \frac{MTE(\lambda | \hat{\lambda})}{\varepsilon^2} = \frac{\text{Var}_{\lambda} \hat{\lambda} + b^2(\lambda | \hat{\lambda})}{\varepsilon^2}$$

$$\xrightarrow{n \rightarrow \infty} 0 \quad \forall \varepsilon > 0 \text{ και } \forall \lambda \in \Theta = \mathbb{R}_+$$

$$\Rightarrow \hat{\lambda} \xrightarrow{P_{\lambda}} \lambda \quad \forall \lambda$$

$$(ε) \frac{\partial}{\partial \lambda} \log f(x) = \frac{\partial}{\partial \lambda} \{ \log 2 + \log \lambda + \log x - \lambda x^2 \} = \frac{1}{\lambda} - x^2$$

$$\Rightarrow I_1(\lambda) = \text{Var}_\lambda \left\{ \frac{\partial}{\partial \lambda} \log f(X, \lambda) \right\} = \text{Var}_\lambda \left\{ \frac{1}{\lambda} - X_1^2 \right\} = \text{Var}_\lambda (X_1^2) = \text{Var}_\lambda Y_1 = \frac{1}{\lambda^2}$$

Οπότε, το κοίλω φράγμα της C-F-R για την διασπορά της αφερόδου $\hat{\lambda}_n$ είναι: $B_n(\lambda) = \frac{1}{I_n(\lambda)} = \frac{1}{n I_1(\lambda)} = \frac{\lambda^2}{n} \quad \forall \lambda$,

ενώ για την διασπορά της $\hat{\lambda}_n$ είναι: $B'_n(\lambda) = \frac{\left(\frac{\partial}{\partial \lambda} E \hat{\lambda}_n \right)^2}{I_n(\lambda)}$
 $= \left(\frac{n}{n-1} \right)^2 \frac{\lambda^2}{n} < \left(\frac{n}{n-1} \right)^2 \frac{\lambda^2}{n-2} = \text{Var}_\lambda \hat{\lambda}_n \quad \forall \lambda$.

Αλλά επίσης: $\text{Var}_\lambda \hat{\lambda}_n = c^2 \text{Var}_\lambda \hat{\lambda}_n = \left(\frac{n-1}{n} \right)^2 \cdot \left(\frac{n}{n-1} \right)^2 \frac{\lambda^2}{n-2} = \frac{\lambda^2}{n-2} > B_n(\lambda) = \frac{\lambda^2}{n} \quad \forall \lambda$.

(3) Οπότε: $e(\lambda | \hat{\lambda}_n) = \frac{B_n(\lambda)}{\text{Var}_\lambda \hat{\lambda}_n} = \frac{\lambda^2/n}{\lambda^2/(n-2)} = \frac{n-2}{n} = 1 - \frac{2}{n} < 1$
 (αλλά "υονά")

και α.ε. $(\lambda | \hat{\lambda}_n) \xrightarrow{\text{ασυμπτωτ. (αφερόδου)}} \lim_{n \rightarrow \infty} \frac{B'_n(\lambda)}{\text{Var}_\lambda \hat{\lambda}_n} = \lim_{n \rightarrow \infty} \frac{\left(\frac{n}{n-1} \right)^2 \lambda^2/n}{\left(\frac{n}{n-1} \right)^2 \frac{\lambda^2}{n-2}} = \lim_{n \rightarrow \infty} \frac{n-2}{n} = \lim_{n \rightarrow \infty} \left(1 - \frac{2}{n} \right) = 1$.

(4) $\hat{\lambda} = \frac{n}{Y} = \frac{1}{\bar{Y}}$, όπου $Y_i = X_i^2 \sim \text{Exp}(\lambda)$, $\mu := \frac{1}{\lambda}$

$\sqrt{n}(\hat{\lambda} - \lambda) = \sqrt{n} \left(\frac{1}{\bar{Y}} - \frac{1}{\mu} \right) = \sqrt{n} (g(\bar{Y}) - g(\mu))$

Αλλά, από ΚΟΘ $\sqrt{n}(\bar{Y} - \mu) \xrightarrow{d} N(0, \text{Var} Y_1) \stackrel{(\text{όπου } g(x) \equiv 1/x)}{=} N(0, \frac{1}{\lambda^2})$,

και άρα, από το Θ. σταθερότητας της διασποράς, έχω

$\sqrt{n}(\hat{\lambda} - \lambda) \xrightarrow{d} N(0, (g'(\mu))^2 \text{Var} Y_1) \stackrel{d}{=} N(0, \lambda^2)$

$\left(-\frac{1}{\mu^2} \right)^2 \cdot \frac{1}{\lambda^2} = \frac{1}{\lambda^4} \cdot \frac{1}{\lambda^2} = \frac{1}{\lambda^6} = \lambda^{-6} = \frac{1}{\lambda^6} = \frac{1}{I_1(\lambda)}$
 $\stackrel{(\text{όπου } \lambda = 1/\mu)}{=} \frac{1}{I_1(\lambda)}$

ο) $\sqrt{n}(\hat{\lambda} - \lambda) \xrightarrow{d} N(0, \lambda^2)$

g? : $\sqrt{n}(g(\hat{\lambda}) - g(\lambda)) \xrightarrow{d} N(0, (g'(\lambda))^2 \lambda^2)$ (από Θ. Έρ. Διασπ.)

$\Rightarrow g'(\lambda) = \pm \frac{1}{\lambda} \Rightarrow g(\lambda) = \pm \log \lambda \quad \forall \lambda$

Αντ., $\sqrt{n}(\log \hat{\lambda} - \log \lambda) \xrightarrow{d} N(0, 1)$. (Με το " - " ισχύει;)